

Analysis of Exchange Rate Volatility in Peru in the Presence of Structural Breaks

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Abstract: The presence of structural breaks in the analysis of volatility in the financial field has been ignored in research done in Latin America. This paper fills this gap by analyzing the behavior of the dollar exchange rate in Peru by evaluating the impact of structural breaks in volatility forecasting. The return behavior was analyzed for the period 05/01/2010 to 09/30/2021. Econometric analysis was used, which consisted of: (1) use of the modified Iterative Cumulative Sum of Squares (ICSS) algorithm to determine the structural breakpoints; (2) estimation of GARCH Models for the subsamples originated by the identified breakpoints; and (3) comparison of alternative models with the GARCH(1,1) expanding window model for horizons of 1, 20, 60 and 120 days. The ICSS algorithm identified 8 breaks in volatility behavior. The models were compared based on out-of-sample forecast performance. It was determined that the GARCH model that considers structural breaks is only effective for a one-day horizon. Finally, the GARCH(1,1) 0.25 rolling window model provides a better strategy for forecasting the volatility of exchange rate returns in Peru for longer horizons.

Keywords: volatility, GARCH, structural breaks, exchange rate.

秘鲁存在结构性中断时的汇率波动分析

摘要: 拉丁美洲的研究忽略了金融领域波动性分析中存在的结构性中断。本文通过评估结构性中断对波动率预测的影响,分析秘鲁的美元汇率行为,填补了这一空白。分析了 2010 年 5 月 1 日至 2021 年 9 月 30 日期间的回报行为。使用计量经济学分析,包括:(1)使用改进的迭代累积平方和算法确定结构断点;(2)对由已识别断点产生的子样本的广义自回归条件异方差模型的估计;(3)替代模型与广义自回归条件异方差(1,1)扩展窗口模型比较,用于 1、20、60 和 120 天的视野。迭代累积平方和算法识别出 8 个波动行为中断。这些模型是根据样本外预测性能进行比较的。已确定考虑结构断裂的广义自回归条件异方差模型仅在一日范围内有效。最后,广义自回归条件异方差(1,1) 0.25 滚动窗口模型为预测秘鲁长期汇率回报的波动性提供了更好的策略。

关键词: 波动率, 广义自回归条件异方差模型, 结构性断裂, 汇率。

1. Introduction

The modeling of the exchange rate volatility process gives solidity to the economic policy structure of a nation. It generates bases for the democratic dynamics necessary for developing a system, but due to the heteroscedastic nature of the short-run exchange rate returns. These classic models assume stationarity and homoscedasticity are not the most suitable for

modeling it. Nevertheless, such modeling is of great importance in macroeconomic decision-making and, in recent decades, has been the subject of research in the financial market.

The exchange rate frequently exhibits periods of large price swings followed by periods of calm. Lately, a macroeconomic analysis has become an important topic of exchange rate movements [1-3]. The issue of

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exchange rate volatility is widely studied due to its potential influence on inflation, international trade, the degree of competitiveness of the economy's external sector, and its role in investment analysis, securities valuation, risk management, and profitability. [4] showed that exchange rate volatility has a significant effect on export products. In addition, exchange rate prediction is considered important for establishing costs and benefits [5].

The behavior of volatility is mathematically expressed from the GARCH. However, these models do not consider changes in volatility, called structural breaks. That induces large biases, giving rise to poor estimates of volatile processes [6]. Structural breaks are relevant in volatility modeling using GARCH. Their inclusion usually improves out-of-sample volatility forecasts [7].

There are studies concerning the volatility in the Latin American environment without considering the presence of breaks [8]. Another research [9] concluded that exchange rate volatility has a significant impact on the export flows of eight Latin American countries.

In the Latin American environment, few studies on volatility consider the presence of breaks [8]. Research that did not consider breaks concludes that exchange rate depreciation has a greater impact on exchange rate volatility [9].

In Peru, [10] analyzed the IGBVL volatility, concluding that a GARCH(1,1) model best explains the variance behavior. In addition, [11] concludes that the behavior of the exchange rate is generally influenced by information from its previous behavior.

No studies analyze the influence of breaks in volatility, both in Peru and in Latin America. For this reason, the objective of this paper is to examine the temporal variation of the volatility of the U.S. dollar exchange rate in Peru during the period 2010-2021, considering the presence of structural breaks. Furthermore, the period under study is considered because it includes recent events that Peru has experienced, both in the political arena and in the health area, which caused instability in the dollar price in Peru.

2. Materials and Methods

2.1. Data and Variables

This study uses the daily dollar exchange rate in Peru for the period 01/05/2010 to 09/30/2021, with a total of 2,922 data. The data were obtained from the Central Bank of Reserves (BCR) website. Daily dollar exchange rate returns were obtained from the exchange rate. The data were processed using free software such as R version 4.1 and Python version 3.9.

2.2. Econometric Methodology

2.2.1. Modified ICSS Algorithm

If we denote the daily return on the dollar as $r_t = 100[\ln(P_t) - \ln(P_{t-1})]$, from time $t - 1$ to period t , for $t = 1, \dots, T$. P_t denotes the selling price of the dollar at time t and $e_t = r_t - \mu$, where μ is the (unconditional) constant mean of r_t .

For checking structural changes in variance occurring at some point in time, the modified ICSS algorithm was applied [12]. The research [13] provides the ICSS algorithm to determine structural changes in the unconditional variance, but this method is not suitable for dependent processes like GARCH models. In this sense, [12] makes a non-parametric adjustment to the algorithm proposed by [13].

2.2.2. In-Sample Test

The In-Sample Test procedure is divided into two phases:

Phase 1: The modified ICSS algorithm determines structural breaks in the unconditional volatility of the series under study.

Phase 2: We estimate GARCH(1,1) models for the complete series and the subsamples defined from the breakpoints identified in phase 1, allowing us to evaluate their empirical relevance.

We express the GARCH(1,1) model for e_t with zero mean as

$$e_t = \sigma_t \varepsilon_t \quad (1)$$

$$\sigma_t^2 = \omega + \alpha e_{t-1}^2 + \beta \sigma_{t-1}^2 \quad (2)$$

with $\omega > 0$, $\alpha \geq 0$, $\beta \geq 0$. If $\alpha + \beta < 1$, the series e_t has a finite variance and being a martingale in differences, it is white noise, of mean zero and variance $Var(e_t) = \frac{\alpha}{1-\alpha-\beta} \cdot \sigma_t^2$ represents the conditional volatility of r_t , and ε_t is i.i.d. with mean zero and unit variance. The model coefficients are often estimated using the quasi-maximum likelihood method (QMLE) due to its estimators' consistency and asymptotic normality [14, 15].

2.2.3. Out-of-Sample Test

We analyzed the out-of-sample volatility forecasts by dividing them into two parts: the first R observations constituted the in-sample component, and the last P observations the out-of-sample component, such that $T=R+P$. We used the GARCH(1,1) expanding window model as the reference model. In addition, to analyze the possible effects coming from structural breaks in the volatility forecast, we used seven forecasting approaches for horizons of 1, 20, 60, and 120 days.

FIGARCH(1,d,1) represents the slow decay of the autocorrelations of the squared returns allowing a significant improvement in their prediction for a long memory process. The model is expressed as

$$\sigma_t^2 = \omega + \beta \sigma_{t-1}^2 + [1 - \beta - (1 - \phi)L - (1 - \phi)L^d] z_t^2 \quad (3)$$

where L is the lag operator.

GARCH(1,1) 0.5 rolling window uses a rolling estimation window equal to half the size of the period in the sample to make the forecast.

GARCH (1,1) 0.25 rolling window uses a moving estimation window equal to one-fourth of the in-sample period size.

GARCH(1,1) expanding window using weighted maximum likelihood estimators. This model handles the structural instability that can appear in GARCH parameters very well.

GARCH(1,1) takes into account the presence of structural breakpoints.

Moving Average using the Average of the squared returns over the previous 300 days to forecast volatility on day t. At longer horizons, the moving average model can improve the daily volatility forecast compared to GARCH(1,1) models, even more so in the presence of structural breaks.

GJR-GARCH(1,1) expanding window is used to capture the effect of leverage.

We use a loss function to compare volatility forecasts between the different models used. The aggregate version of the proposed metric is the mean squared forecast error (MSFE) which is given by

$$MSFE_{S,1} = [P - (S - 1)]^{-1} \sum_{t=R+2}^T (\tilde{e}_t^2 - \tilde{\sigma}_{t|R-s,1}^2)^2 \quad (4)$$

where

$$\tilde{e}_t^2 = \sum_{j=1}^s e_{t-(j-1)}^2 y \quad \tilde{\sigma}_{t|R-s,i}^2 = \sum_{j=1}^s \sigma_{t-(j-1)}^2 | t - s, i. \quad (5)$$

Significant differences in MSFE between the reference model and a competition model were tested using the statistic developed by [16]. This statistic allows us to establish significant differences in the forecasts of nested models. Therefore, the competing models in this research are considered nested models.

3. Results

3.1. Data and Descriptive Statistics

We analyzed the behavior of the daily return of the dollar in Peru in the period 05/01/2010 to 09/30/2021. Descriptive statistics are reported in Table 1. We use the closing prices to calculate the daily dollar returns. The mean return is 0.0123 with a standard deviation of 0.2935.

An excess Kurtosis (14.2293) also indicates a leptokurtic distribution, which is an apparent deviation from normality. The daily stock returns appear very volatile. Furthermore, the presence of serial correlations in the squared returns and the confirmation of ARCH effects ($p < 0.01$) support the use of GARCH models to describe the behavior of the volatility of dollar price returns in Peru.

Table 1 Summary statistics of exchange rate returns

	Returns	p-value
Return Average	0.0123	
Standard Deviation	0.2935	
Asymmetry	-0.2480	
Kurtosis	14.2293	
Minimum	-2.9638	
Maximum	3.0123	
Square Exchange Rate Return Ljung-Box (r = 20)	486.6821	0.000
ARCH LM (q = 2)	200.7360	0.000
ARCH LM (q = 10)	278.446	0.000

Fig. 1 shows the ups and downs in the exchange rate during the sample period. It shows periods in which large changes are followed by other large changes and periods in which other small changes follow small changes. In addition, the significant point to note is that volatility occurs in clusters.

Periods of relative tranquility are seen, especially between 2014 - 2015 and 2017 - 2019, during which small positive and negative returns were present. On the other hand, during the years 2016, 2020, and part of 2021, there was greater volatility, observing positive and negative returns of greater magnitude. This previous analysis concludes that volatility is autocorrelated [17].

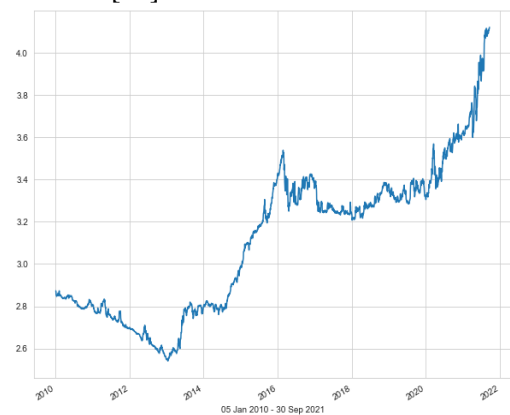


Fig. 1 Dollar selling price

3.2. In-Sample Results

Table 2 shows the eight break points in conditional volatility. The breakpoints are also presented in Fig. 2 and were found using the modified ICSS algorithm.

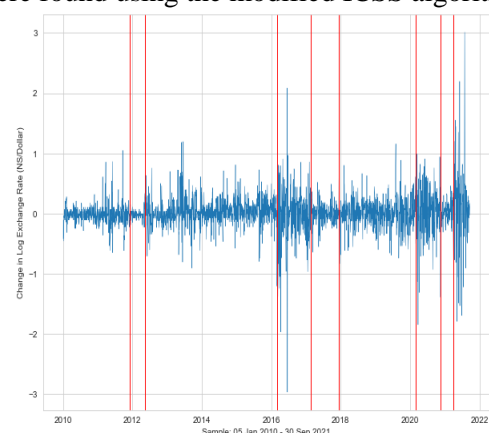


Fig. 2 Structural breakpoints

Table 3 shows the estimated coefficients of the GARCH(1,1) models for the total sample period and the subsamples. For the total sample, in the fitted GARCH(1,1) model, we find that $\alpha + \beta \approx 0.999$ indicates a high volatility persistence. The subsamples GARCH(1,1) models, except for subsamples 2, 6, and 8, present estimates of $\alpha + \beta > 0.95$ indicating high persistence and characterizing them by conditional heteroscedasticity. According to [7], these changes in the persistence of the volatility process may be due to not considering structural breaks.

In addition, the unconditional variance presented significant changes, which are reflected by the values of $\omega / (1 - \alpha - \beta)$. These changes are due to breakpoints that cause a significant change in ω . In addition, the values of the estimated coefficients of the

GARCH(1,1) model vary between the subsamples. Overall, the results indicate that the conditional variance of exchange rate returns is highly persistent and that structural breaks characterize them significantly.

Table 2 Structural breaks in volatility with a time period

Sample	Structural breaks	Time Period
Total Sample	8	05/01/2010 - 20/09/2021
Subsample 1		05/01/2010 - 01/12/2011
Subsample 2		02/12/2011 - 11/05/2012
Subsample 3		11/05/2012 - 01/03/2016
Subsample 4		02/03/2016 - 21/02/2017
Subsample 5		22/02/2017 - 14/12/2017
Subsample 6		15/12/2017 - 27/02/2020
Subsample 7		28/02/2020 - 19/11/2020
Subsample 8		20/11/2020 - 05/04/2021
Subsample 9		05/04/2021 - 30/09/2021

Table 3 QMLE estimation results for GARCH models (1,1)

	Total Sample	Subsample 1	Subsample 2	Subsample 3
ω	0.001 (0.00)	0.001 (0.001)	0.002 (0.009)	0.003 (0.001)
α	0.234 (0.04)	0.672 (0.23)	0.835 (0.46)	0.162 (0.05)
β	0.765 (0.04)	0.321 (0.13)	0.005 (0.33)	0.776 (0.06)
ω	1.56	0.06	0.012	0.05
$1 - \alpha - \beta$				
	Subsample 4	Subsample 5	Subsample 6	Subsample 7
ω	0.020 (0.01)	0.003 (0.002)	0.015 (0.008)	0.009 (0.01)
α	0.409 (0.14)	0.201 (0.10)	0.224 (0.10)	0.190 (0.18)
β	0.556 (0.08)	0.714 (0.010)	0.525 (0.18)	0.783 (0.12)
ω	0.57	0.03	0.06	0.33
$1 - \alpha - \beta$				
	Subsample 8	Subsample 9		
ω	0.010 (0.01)	0.000 (0.01)		
α	0.205 (0.19)	0.110 (0.07)		
β	0.434 (0.19)	0.889 (0.07)		
ω	0.028	0.00		
$1 - \alpha - \beta$				

3.3. Out-of-Sample Results

We consider the last 300 observations of the entire sample period and cover the period from July 23, 2020, to September 30, 2021, including the last two structural breaks. Table 4 presents the out-of-sample volatility forecast results for the different horizons. The MSFE for each forecasting method and its comparison ratio with the MSFE of the GARCH(1,1) expanding window model are presented. A model with a ratio less than unity performs better than the model in reference. The test developed by [15] was applied to determine whether the competing model, compared to the GARCH(1,1) expanding window model, had a lower expected loss.

We compared the GARCH(1,1) expanding window model in the first two rows of each horizon and the FIGARCH(1,d,1) model. The benchmark model performs better at horizons $s = 1$ and 20, and the FIGARCH model performs better at horizons $s = 60$ y 120, indicating that this model is suitable for long memory processes.

On the other hand, the GARCH(1,1) model with breaks shows a lower average loss rate at the 1-day

horizon than the benchmark and competing models. However, its performance worsens at higher forecast horizons, although it performs relatively better than the GJR-GARCH(1,1) expanding window and moving average models. At the 60 and 120-day horizons, the GARCH(1,1) 0.25 rolling window model performs better than the benchmark model. Moreover, as mentioned above, the GARCH(1,1) model with breaks cannot outperform the benchmark model. These results allow us to conclude that it is not easy to determine an estimation window that would allow us to make more accurate volatility forecasts from a GARCH(1,1) model. Weighted maximum likelihood and moving average GARCH(1,1) models never offer volatility forecast gain related to the benchmark.

Table 4 Out-of-sample prediction results

	s = 1		s = 20	
	MSFE	Ratio	MSFE	Ratio
GARCH(1,1) expanding window	2.50	1.00	61.98	1.00
FIGARCH expanding window	2.65	1.06	102.34	1.65
GARCH(1,1) 0.50 rolling window	4.70	1.88	164.24	2.65

Continuation of Table 4				
GARCH(1,1) 0.25 rolling window	6.33	2.53	182.73	2.95
GARCH(1,1) MV weighted	6.56	2.62	185.44	2.99
GARCH(1,1) with breaks	2.39	0.96*	63.31	1.02
Moving Average	3.17	1.27	199.22	3.21
GJR-GARCH(1,1) expanding window	4.48	1.79	55.07	0.89*
	s = 60		s = 120	
	MSFE	Ratio	MSFE	Ratio
GARCH(1,1) expanding window	536.90	1.00	2583.68	1.00
FIGARCH expanding window	514.66	0.96	2161.92	0.84
GARCH(1,1) 0.50 rolling window	1026.47	1.91	3162.20	1.22
GARCH(1,1) 0.25 rolling window	415.32	0.77*	2002.51	0.77*
GARCH(1,1) MV weighted	581.92	1.08	3302.74	1.28
GARCH(1,1) with breaks	557.41	1.07	2830.25	1.09
Moving Average	1075.23	2.00	3102.23	1.20
GJR-GARCH(1,1) expanding window	1070.11	1.99	3967.76	1.54

* Significant ratio at $p < 0.05$

4. Discussion

A proper analysis of volatility in financial series, such as stock market shares and the exchange rate, is crucial in macroeconomic decision-making. Its accurate modeling allows for establishing a sound economic policy of a country [17]. However, this volatile behavior of financial series is not constant over time and usually presents changes in the mathematical structure that tries to explain its behavior.

Much research on forecasting exchange rate volatility, but unfortunately, the presence of structural breaks has not been considered in previous research on exchange rate volatility in Latin America, especially in Peru. [11] did not consider the presence of structural breaks, assuming a stable GARCH process.

The present paper shows that the GARCH(1,1) 0.25 rolling window model for longer horizons shows a significant gain. However, as expressed by [17], the optimal size of the estimation window used in the prediction model is related to the rupture's size, moment, and direction, which makes its selection difficult.

5. Conclusions

Although GARCH models are frequently used in volatility forecasting, little attention has been paid to the impact of structural breaks on such forecasts. Research assumes the existence of a stable volatility process. It, therefore, uses a recursive (expanding) or fixed window size when estimating the GARCH model used to generate out-of-sample volatility forecasts. If there is a structural break in the volatility pattern, the

model's volatility forecast may no longer be consistent or reliable without considering the structural break. As a result, GARCH models do not accurately track changes in unconditional variance, resulting in forecasts that underestimate or overestimate volatility over long periods. To this, we add that in Latin America, there are no studies that consider the presence of structural breaks in the analysis of volatility.

Our study differs in several aspects from previous studies in Latin America. First, we consider the presence of structural breaks in volatility analysis by applying an econometric methodology that allows us to identify them and not consider a stable GARCH process in volatility forecasting. Second, it focuses on an important emerging market in Latin America: Peru, for which the existing literature indicates a gap in this type of analysis. Finally, we used the modified ICSS algorithm to identify breakpoints and estimated GARCH models for each identified subsample. That allowed us to combine forecasts of estimated volatility models and use different window sizes, which provides a better strategy for forecasting exchange rate volatility.

The results show the relevance of structural breaks in the GARCH models when explaining the volatility behavior of the U.S. dollar exchange rate performance in Peru. Applying the modified ICSS algorithm in the sample found eight structural breaks in the GARCH(1,1) processes, resulting in significant changes in the GARCH(1,1) parameter estimates across subsamples.

Using an out-of-sample period from July 23, 2020, to September 30, 2021, for Peru, including the last two structural breaks, we find the GARCH(1,1) model with structural breaks only shows a significant gain in predictive ability compared to the benchmark GARCH(1,1). That expands the window model at a horizon equal to one. The GJR-GARCH(1,1) expanding window model, estimated to capture the effects of leverage, showed a significant gain in success at a horizon equal to 20. On the other hand, the GARCH(1,1) 0.25 rolling window model shows significant gain at horizons equal to 60 and 120.

In addition, the combination of forecasting models using different windows provides a better strategy for forecasting the volatility of exchange rate returns in Peru.

The study is one of the first to evidence the presence of structural breaks in the volatile behavior of an economic variable in Latin America, especially in Peru. Therefore, we hope that regulators and financial policymakers will consider this study for the creation of policies that support the stock market. That is because the exchange rate is an important factor that investors should consider before investing in stocks.

References

- [1] MUMTAZ A. Impact of exchange rate and its volatility on domestic consumption in India and Pakistan. *Journal of Public Affairs*, 2020: e2479. <https://doi.org/10.1002/pa.2479>
- [2] ASTERIOU D., MASATCI K., and PIPEAM K. Exchange rate volatility and international trade: International evidence from the MINT countries. *Economic Modelling*, 2016, 59(133). <https://doi.org/10.1016/j.econmod.2016.05.006>
- [3] NOR M., MASRON T., and ALABDULLAH T. Macroeconomic fundamentals and the exchange rate volatility: Empirical evidence from Somalia. *SAGE Open*, 2020, 10(1). <https://doi.org/10.1177/2158244019898841>
- [4] SUGIHARTI L., ESQUIVIAS M. A., and SETYORANI B. The impact of exchange rate volatility on Indonesia's top exports to the five main export markets. *Heliyon*, 2020, 6(1). <https://doi.org/10.1016/j.heliyon.2019.e03141>
- [5] EPAPHRA M. Modeling Exchange Rate Volatility: Application of the GARCH and EGARCH Models. *Journal of Mathematical Finance*, 2017, 7: 121-143. <https://doi.org/10.4236/jmf.2017.71007>
- [6] TRENCA I., ZAPODEANU D., and COCIUBA M. Testing the Presence of Structural Break in the Euro Exchange Rate Variance. *Procedia Economics and Finance*, 2015, 32: 1163-1169. [https://doi.org/10.1016/S2212-5671\(15\)01582-8](https://doi.org/10.1016/S2212-5671(15)01582-8)
- [7] ŽIVKOV D., NJEGIĆ J., and MOMČILOVIĆ M. Bidirectional spillover effect between Russian stock index and the selected commodities. *Zbornik Radova Ekonomskog Fakulteta au Rijeci*, 2018, 36: 27-51. <https://doi.org/10.18045/zbefri.2018.1.29>
- [8] GÜLOGLU B., KAYA P., and AYDEMİR R. Volatility Transmission among Latin American Stock Markets under Structural Breaks. *Physica A*, 2016, 462: 330-340. <https://doi.org/10.1016/j.physa.2016.06.093>
- [9] ROJAS E., GUZMAN J., and BALTAZAR J. Volatilidad cambiaria, metas de inflación y crisis financiera global. Evidencia para economías latinoamericanas. *Revista Economía y Política*, 2019, 15(30). <https://doi.org/10.25097/rep.n30.2019.07>
- [10] SOTOMAYOR R. N., & CASTILLO J. E. Modelamiento de la volatilidad del Índice General de la Bolsa de Valores de Lima, periodo 2009-2011. *Anales Científicos*, 2016, 77(1): 1-7. <https://doi.org/10.21704/ac.v77i1.576>
- [11] CHUNG V. Modelación de la Volatilidad del Tipo de Cambio del Dólar en el Perú: Aplicación de los Modelos GARCH y EGARCH. *Análisis Económico y Financiero*, 2021, 4(2): 7-12. <https://doi.org/10.24265/raef.2021.v4n2.40>
- [12] JINLAN M., & BEICHEN W. Structural Breaks in Volatility: The Case of Chinese Stock Returns. *The Chinese Economy*, 2016, 49(2): 81-93. <http://dx.doi.org/10.1080/10971475.2016.1143302>
- [13] INCLÁN C., & TIAO C. Use of cumulative squares for retrospective detection of changes in variance. *Journal of the American Statistical Association*, 1994, 89: 913-923. <https://doi.org/10.1080/01621459.1994.10476824>
- [14] BLASQUES F., GORGI P., and SIEM J. Feasible invertibility conditions and maximum likelihood estimation for observation-driven models. *Electronic Journal of Statistic*, 2018, 12: 1019-1052. <https://doi.org/10.1214/18-EJS1416>
- [15] STARICA C., & GRANGER C. Nonstationarities in

Stock Returns. *The Review of Economics and Statistics*, 2005, 87(3): 503-522. <https://www.jstor.org/stable/40042945>

[16] BROOKS C. *Introductory Econometrics for Finance*. Cambridge University Press, Cambridge, 2019.

[17] BABIKIR A., & GUPTA R. Structural breaks and GARCH models of stock return volatility: The case of South Africa. *Economic Modelling*, 2012, 29(6): 2435-2443. <https://doi.org/10.1016/j.econmod.2012.06.038>

参考文献:

[1] MUMTAZ A. 汇率及其波动对印度和巴基斯坦国内消费的影响。公共事务杂志, 2020: e2479. <https://doi.org/10.1002/pa.2479>

[2] ASTERIOU D., MASATCI K., 和 PIPEAM K. 汇率波动和国际贸易: 来自墨西哥、印度尼西亚、尼日利亚和土耳其的国际证据。经济建模, 2016, 59(133). <https://doi.org/10.1016/j.econmod.2016.05.006>

[3] NOR M., MASRON T., 和 ALABDULLAH T. 宏观经济基本面和汇率波动: 来自索马里的经验证据。圣公开赛, 2020, 10(1). <https://doi.org/10.1177/2158244019898841>

[4] SUGIHARTI L., ESQUIVIAS M. A., 和 SETYORANI B. 汇率波动对印尼对五个主要出口市场的主要出口的影响。赫利昂, 2020, 6(1). <https://doi.org/10.1016/j.heliyon.2019.e03141>

[5] EPAPHRA M. 建模汇率波动性: 广义自回归条件异方差和指数广义自回归条件异方差模型的应用。数学金融杂志, 2017, 7: 121-143. <https://doi.org/10.4236/jmf.2017.71007>

[6] TRENCA I., ZAPODEANU D., 和 COCIUBA M. 测试欧元汇率差异中是否存在结构性突破。普罗西迪亚经济与金融, 2015, 32: 1163-1169. [https://doi.org/10.1016/S2212-5671\(15\)01582-8](https://doi.org/10.1016/S2212-5671(15)01582-8)

[7] ŽIVKOV D., NJEGIĆ J., 和 MOMČILOVIĆ M. 俄罗斯股票指数与选定商品之间的双向溢出效应。里耶卡经济学院学报, 2018, 36: 27-51. <https://doi.org/10.18045/zbefri.2018.1.29>

[8] GÜLOGLU B., KAYA P., 和 AYDEMİR R. 结构性中断下拉丁美洲股市的波动性传导。物理学甲, 2016, 462: 330-340. <https://doi.org/10.1016/j.physa.2016.06.093>

[9] ROJAS E., GUZMAN J., 和 BALTAZAR J. 汇率波动、通胀目标和全球金融危机。拉丁美洲经济的证据。经济与政治杂志, 2019, 15(30). <https://doi.org/10.25097/rep.n30.2019.07>

[10] SOTOMAYOR R. N., 和 CASTILLO J. E. 利马证券交易所综合指数波动率建模, 2009-2011 年期间。科学年鉴, 2016, 77(1): 1-7. <https://doi.org/10.21704/ac.v77i1.576>

[11] CHUNG V. 秘鲁美元汇率波动率的建模: 广义自回归条件异方差模型和指数广义自回归条件异方差模型的应用。经济和金融分析, 2021, 4(2): 7-12. <https://doi.org/10.24265/raef.2021.v4n2.40>

[12] JINLAN M., 和 BEICHEN W. 波动性的结构性突破: 以中国股票回报为例。中国经济, 2016, 49(2): 81-93. <http://dx.doi.org/10.1080/10971475.2016.1143302>

[13] INCLÁN C., 和 TIAO C. 使用累积平方来回顾性检测方差的变化。美国统计协会杂志, 1994, 89: 913-923.

<https://doi.org/10.1080/01621459.1994.10476824>

[14] BLASQUES F., GORGI P., 和 SIEM J. 观测驱动模型的可行可逆性条件和最大似然估计。电子统计杂志, 2018, 12: 1019–1052. <https://doi.org/10.1214/18-EJS1416>

[15] STARICA C., 和 GRANGER C. 股票收益的非平稳性。经济学与统计学评论, 2005, 87(3): 503–522. <https://www.jstor.org/stable/40042945>

[16] BROOKS C. 金融计量经济学导论。剑桥大学出版社, 剑桥, 2019.

[17] BABIKIR A., 和 GUPTA R. 股票收益波动的结构性中断和广义自回归条件异方差模型：南非案例。经济建模, 2012, 29(6): 2435–2443. <https://doi.org/10.1016/j.econmod.2012.06.038>